

1 Rapid speciation in small populations challenges 2 the dominance of ecological speciation

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10 **Abstract**

11 Speciation – the process by which two lineages become reproductively iso-
12 lated – plays a key role in the emergence and maintenance of biodiversity. Yet, our
13 understanding of the time it takes for speciation to occur, and of the microevolu-
14 tionary processes that influence this tempo, remains limited. Here, we thoroughly
15 characterize how population size, mutation rate, local adaptation and migration are
16 expected to influence the duration of speciation, as well as the shape of the “grey
17 zone” of speciation. We show that the relationship between population size and
18 speciation time is indicative of the speciation mode, as faster speciation in smaller
19 populations only occurs in the case of non-ecological speciation. Leveraging ge-
20 nomic estimates of population size and speciation duration across 196 pairs of
21 plant species, we uncover a positive association between population size and spe-
22 ciation duration. Taken together, these results challenge the view that ecological
23 speciation is the source of much of species diversity.

24 **Significance statement**

25 How new species arise, and how quickly they do so, plays a key role in shaping
26 Earth’s biodiversity. Yet, the links between how speciation occurs, how fast it
27 proceeds, and basic properties of species such as population size, remain debated.
28 We clarify these relationships using an integrative model of speciation. We show,
29 in particular, that faster speciation in smaller populations is a signature of non-
30 ecological speciation. Analyses of genomic data from plants reveal precisely this
31 pattern, challenging the prevailing view that speciation in nature is predominantly
32 ecological.

33 1 Introduction

34 Speciation – the process by which a new species emerge – is one of the most fundamen-
35 tal processes in biology. It is at the origin of species diversity. In spite of substantial
36 research on speciation, many controversies remain on its mode, i.e, how it occurs, and
37 tempo, i.e., at which pace it occurs (Coyne and Orr, 2004). There is however a general
38 consensus that speciation takes time (Benton and Pearson, 2001; Etienne et al., 2014),
39 except in exceptional cases, such as speciation by polyploidization in plants (Rieseberg
40 and Willis, 2007) or speciation by host shifts in pathogens (Giraud et al., 2010). In-
41 tuitively, fast-speciating groups will experience speciation events more frequently than
42 slow-speciating ones. Together with the frequency of extinction events, these differ-
43 ences can explain differences in species richness across groups, and together with the
44 frequency of dispersal events, differences in species richness across geographic regions
45 (Wiens and Donoghue, 2004; Schluter and Pennell, 2017).

46 Insights on the tempo of speciation can be gained by fitting diversification models
47 to extant phylogenies (Nee, 2006; Morlon et al., 2024). Fitting birth-death models
48 representing lineages birth (speciation) and death (extinction), assumed to occur in-
49 stantaneously, provides estimates of speciation and extinction rates, i.e., estimates of
50 the average number of events occurring per lineage in a given amount of time. Such
51 studies have shown that speciation rates vary by several orders of magnitude across
52 lineages (Rabosky, 2016; Maliet et al., 2019; Quintero et al., 2024). For example, esti-
53 mates range from 0.01 to 5 spp · Myr⁻¹ in birds (Maliet et al., 2019), and from 0.005 to
54 1.5 spp · Myr⁻¹ in mammals (Quintero et al., 2024).

55 A central challenge lies in understanding if and how microevolutionary (intraspecific)
56 processes, such as drift, selection and migration, can explain this speciation rate hetero-
57 geneity (Harvey et al., 2019; Rolland et al., 2023; Morlon et al., 2024). In comparison
58 with the numerous studies that have investigated correlates of speciation rates with
59 species traits, as well as the abiotic and biotic environment they experience (Benton
60 and Pearson, 2001; Lagomarsino et al., 2016; Schluter and Pennell, 2017; Landis et al.,
61 2022; Wiens, 2024), empirical studies correlating speciation rates to metrics reflecting
62 microevolutionary processes remain rare. The few studies that have investigated the
63 relationship between speciation rates and such parameters have found contrasting or
64 counter-intuitive results, in particular in terms of correlations with the velocity of re-
65 productive isolation (Rabosky and Matute, 2013), population divergence rates (Harvey
66 et al., 2017, 2019), population structure (Singhal et al., 2018; Burbrink et al., 2023),
67 substitution rates (Lanfear et al., 2010; Goldie et al., 2011), or genetic diversity (Huang
68 et al., 2018; Perez-Lamarque et al., 2022; Afonso Silva et al., 2025).

69 Mathematical models have been central in our understanding of the dynamics of speci-
70 ation, and of the microevolutionary processes that influence these dynamics (Gavrilets,
71 2003; Coyne and Orr, 2004). In particular, Gavrilets’ holey adaptive landscape (HAL)
72 provided a mathematically tractable extension of Wright’s rugged adaptive landscape
73 (Wright, 1932) to high dimensional genotypic spaces (Gavrilets et al., 1998). He
74 showed that – simply as a result of the high dimensionality of adaptive landscapes –
75 populations can become separated by regions of low fitness (“holes”), and thus become

76 reproductively isolated, without crossing deep valleys of low fitness by evolving along
77 nearly neutral ridges (Gavrilets, 2003). This framework is therefore well suited for
78 analyzing how drift and selection jointly influence the dynamics of speciation. In his
79 1999 paper, Gavrilets derived a series of mathematical results describing the dynamics
80 of divergence within- and between-populations under strict allopatry, in the presence
81 of migration, and with local adaptation. These provide the foundation for obtaining an-
82 alytical results concerning the expected duration of speciation, although those results
83 were not derived. A later review focused on this question (Gavrilets, 2003) but with-
84 out taking into account within-population genetic variation. Yet, polymorphism can
85 substantially alter the speed at which mutations causing reproductive incompatibilities
86 accumulate, and therefore speciation duration (Gavrilets, 1999).

87 Population size has occupied a central position in debates about the relative role of drift
88 and selection in the speciation process (Coyne and Orr, 2004). This idea traces back to
89 Mayr’s verbal “genetic revolution” hypothesis that population bottlenecks, for exam-
90 ple during founder events, can trigger a shift in allele frequency over several linked loci
91 leading to speciation (Mayr, 1963). Under this hypothesis, speciation should be faster
92 in smaller populations. Opponents of this hypothesis, however, predicted that specia-
93 tion should be faster in larger populations, where natural selection is more efficient
94 (Orr and Orr, 1996). Recent developments in phylogenetic models of diversification,
95 combined with the availability of large empirical datasets, have allowed for new tests
96 of the relationship between proxies of population size, such as range size or nucleotide
97 diversity, and speciation rates (Maya-Lastra and Eaton, 2021; Smyčka et al., 2023;
98 Afonso Silva et al., 2025). These tests, however, treat speciation as an instantaneous
99 event, which rate is only weakly influenced by the duration of speciation (Veron
100 et al., 2025), and can be confounded by the reciprocal effect of speciation on population
101 sizes (Smyčka et al., 2023). Inferences from speciation genomics, on the other hand,
102 can provide estimates of speciation duration *per se*, along with estimates of population
103 sizes before and after the speciation event (Fraïsse et al., 2021).

104 Here, we begin by deriving new analytical predictions for the duration of speciation
105 under the joint influence of drift, selection and gene flow, focusing on the effect of
106 population size. Our derivations build upon Gavrilet’s holey adaptive landscape (HAL)
107 model (Gavrilets, 1999). We then design a test based on our theoretical results on the
108 relationship between speciation duration and population size to assess speciation mode
109 based on genomic data, which we apply to 196 pairs of plant species (Monnet et al.,
110 2025).

111 2 Results and discussion

112 We consider the HAL model of speciation (Gavrilets et al., 1998; Gavrilets, 1999;
113 [Materials and methods](#)). Initially, mutations arise in a population made of N individ-
114 uals, with a per-individual per-generation mutation rate ν . Any pair of individuals is
115 thus characterized by a genetic distance d , defined as the number of nucleotide sites
116 at which they differ. If d is above a given threshold K called “genetic incompatibility
117 threshold”, the individuals cannot interbreed. Hence K is the maximum number of loci

118 on which two individuals can be different and still be interfertile. This mathematical
 119 simplification is motivated by the decrease in the fitness of hybrids with the genetic
 120 distance between their parents (Price and Bouvier, 2002; Rabosky and Matute, 2013;
 121 Christie and Strauss, 2018; Christie et al., 2022). This framework provides a mathe-
 122 matically tractable way to represent high dimensional, correlated adaptive landscapes
 123 (Gavrilets, 1997, figure 1A), and to account for the multi-genic origin of incompati-
 124 bilities without the complexity of a large set of Dobzhansky-Muller incompatibilities
 125 (DMIs; Maya-Lastra and Eaton, 2021). Even if mutations are neutral in themselves,
 126 they eventually lead to the reproductive incompatibility of genetically distant individ-
 127 uals, as in DMIs. Following Gavrilets (1999), we consider three versions of the model.
 128 In the “allopatric neutral” scenario, the initial population experiences a random split at
 129 time t_{split} , for example due to the emergence of a geographic barrier isolating two pop-
 130 ulations, after which strict isolation is maintained. The two populations then gradually
 131 accumulate divergence, until speciation occurs (figure 1B). In the “parapatric neutral”
 132 scenario, the two populations can exchange alleles through migration, with a migration
 133 rate m . In the “adaptive” scenario, mutations confer a selective advantage in the local
 134 population in which they arise, with coefficient s_{LA} .

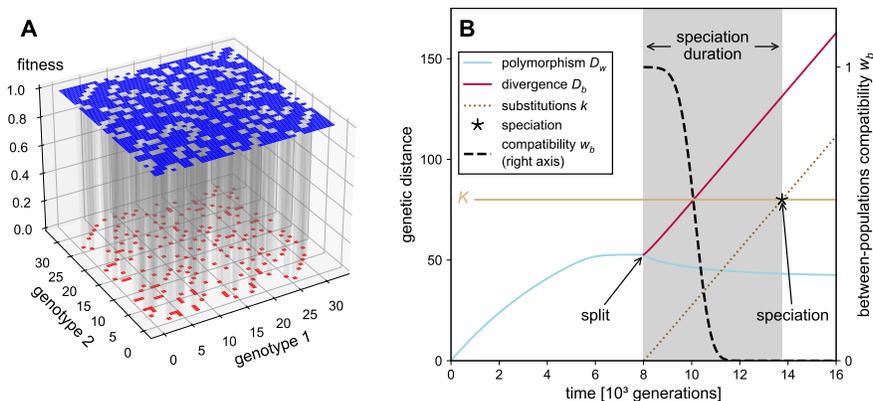


Figure 1. **A.** Illustration of the holey adaptive landscape (HAL) with 5 polymorphic sites (32 different genotypes) and a compatibility threshold $K = 3$. Populations can evolve along nearly neutral ridges (in blue) and end up being separated by fitness holes (in red). **B.** Example of temporal dynamics under the HAL model with neutral mutations and no migration. The ancestral population is split in two after 8000 generations. We track within-population polymorphism (D_w), between-population divergence (D_b), the number of different substitutions between the two populations (k), as well as between population compatibility (w_b , right axis). The star indicates speciation, characterized by the time when $w_b = 0$, or, alternatively, $k(t) = K$ with K the incompatibility threshold. Here $N_T = 6000$, $N_1 = N_2 = 3000$, $\nu = 0.007$, $K = 80$.

135 Following Gavrilets (1999), we use a deterministic approximation to investigate the
 136 temporal dynamics of both within-population polymorphism D_w , defined as the mean
 137 pairwise distance between individuals within populations (w for within), and between-
 138 population divergence D_b , defined as the mean pairwise distance between individuals

139 from the two different populations (b for between; [figure 1, Materials and methods](#)). We
140 also track the probabilities w_w and w_b that two randomly chosen individuals from the
141 same or different populations, respectively, can successfully interbreed. These prob-
142 abilities are referred to as within- and between-population compatibility. Finally, we
143 follow the number k of substitutions (fixed alleles) that differentiate the two popula-
144 tions. We consider speciation to occur if (and when) no individual from one population
145 can successfully interbreed with any individual from the other population (i.e., $w_b = 0$;
146 [Mayr, 1942](#)). This occurs if (and when) k reaches K ([figure 1B](#)); indeed, in this case,
147 any two individuals picked in different populations are distant by at least K loci and
148 are thus incompatible. We also investigate the shape of the grey zone of speciation,
149 when the reproductive compatibility between the populations w_b is reduced or even
150 close to zero, but does not reach zero. This corresponds to the often-encountered inter-
151 mediate situation in nature where populations are not fully compatible, but speciation
152 is not complete. We validate our theoretical predictions using intensive simulations
153 ([supplementary text S2](#) and [figures S7 to S10](#)).

154 In the absence of migration, populations ineluctably diverge and speciation occurs after
155 enough time elapses. In the allopatric neutral scenario, we show that the speciation
156 duration τ , defined as the expected time between t_{split} and the time when speciation
157 occurs, is given by :

$$158 \quad \tau = \frac{K}{\nu(R_1 + R_2)} \quad (1)$$

159 where R_1 and R_2 are, in each of the two sub-populations, coefficients of fixation which
160 capture the efficiency of purifying selection to remove incompatibilities and depend
161 on population size, mutation rate, and the incompatibility threshold ([equation 3](#) in the
162 [Materials and methods](#)). $R \approx 1$ represents the case when this effect is negligible. For
163 simplicity, we assume in what follows that the ancestral population is split into two
164 populations of equal sizes, but our equations can accommodate asymmetric population
165 sizes. As K decreases, and N or ν increase, the effect of purifying selection against in-
166 compatibilities is stronger, and R decreases (see [figure S1](#)), reducing within-population
167 polymorphism ([figure 2](#), lower row) and affecting the duration of speciation ([figure 2](#),
168 upper row, blue curves). Equation 1 holds in the adaptive scenario, with an expression
169 for R that depends on the local selective advantage of mutations s_{LA} ([Materials and](#)
170 [methods](#)). As s_{LA} increases, the mutations in each environment confer a local adaptive
171 advantage, and their fixation is faster. Hence the coefficient of fixation increases and
172 can be larger than 1. As expected, local adaptation speeds up speciation ([figure 2](#), upper
173 row, green curves). In the presence of migration, we can determine when speciation
174 occurs by solving a set of ordinary differential equations ([Materials and methods](#)). As
175 expected, gene flow slows down speciation ([figure 2](#)). In fact, speciation is no longer
176 ineluctable, and we designed an efficient way to predict whether it will occur or not
177 ([Materials and methods](#) and [supplementary text S1](#); [figure 2](#) upper row, pink curves).

178 High mutation rates are generally thought to increase the rate at which populations ac-
179 quire substitutions and reproductive isolation, resulting in higher speciation rates ([Lan-
180 fear et al., 2010](#); [Hua and Bromham, 2017](#)). This expectation is part of the integrated
181 evolutionary speed hypothesis, which stipulates that the short generation times and high

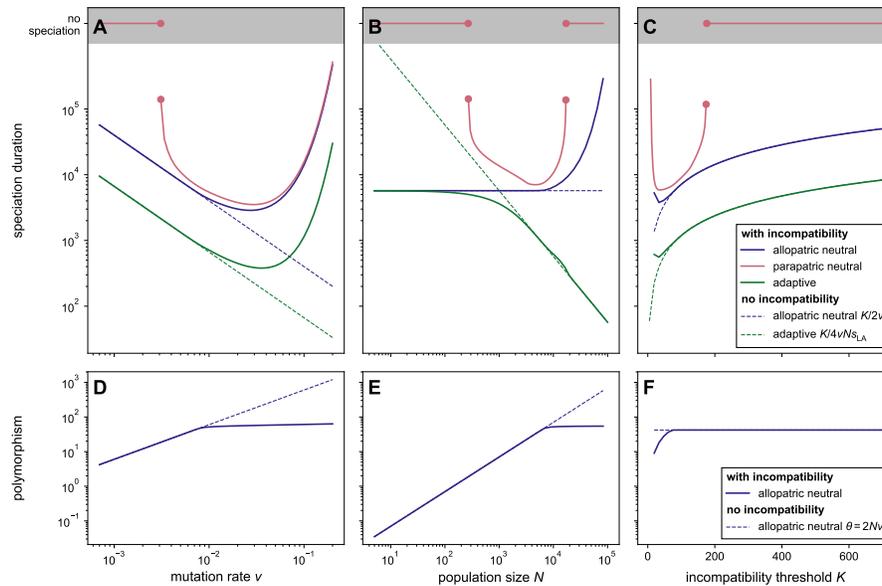


Figure 2. Expected speciation duration (upper row) and polymorphism (lower row) under the holey adaptive landscape model (HAL; solid lines), as a function of mutation rate (A, D), population size (B, E) and incompatibility threshold (C, F). Three versions of the model are considered: strict allopatry (in blue), with migration (parapatric, in red), and with local adaptation (in green). Expectations under a purely theoretical scenario without purifying selection against incompatibilities within populations ($R = 1$, $s = 0$) are shown with dashed lines for comparison. On the upper row, the pink line in the shaded region indicates cases where speciation does not occur. In each panel, one of the three parameters varies with the others kept constant at $\nu = 0.0007$, $N = 6000$, and $K = 80$. For the parapatric scenario, the migration rate is $m = 10^{-4}$; for the adaptive scenario, the coefficient of selection is $s_{LA} = 0.001$.

182 mutation rates found in warm (tropical) environments result in fast genetic change and
 183 speciation (Rohde, 1992). In a purely theoretical scenario without within-population
 184 purifying selection against incompatibilities, increased mutation rates indeed always
 185 speed up speciation (figure 2A, dashed curves). When within-population incompatibil-
 186 ities are selected against, however, this expectation is true only under low-to-moderate
 187 mutation rates; when mutations become more frequent (> 0.05 per generation), pu-
 188 rifying selection against incompatibilities limits their accumulation both within and
 189 between populations, and slows down speciation, contrary to the evolutionary speed
 190 hypothesis (figure 2A, solid curves). This speciation slowdown under a high mutation
 191 rate regime occurs in the three scenarios we considered. The non-monotonic depen-
 192 dency of speciation duration to mutation rates may explain why tests of the evolu-
 193 tionary speed hypothesis have found mixed evidence. A positive association was found
 194 in birds (Lanfear et al., 2010) and plants (Bromham et al., 2015), but a negative one
 195 was found in mammals (Afonso Silva et al., 2025). In the latter study, the authors at-
 196 tributed this *a priori* unexpected result to methodological artifacts; our results suggest

197 that it could in fact be real, due to a high purifying selection, and slow accumulation of
198 incompatibilities, when mutations are frequent.

199 The dependency of speciation duration to population size strongly depends on whether
200 speciation is ecological (adaptive scenario, in green) or not (neutral scenarios, in blue
201 and red) (figure 2B). As expected, ecological speciation is faster in large populations
202 where positive selection is more effective; this effect becomes visible when the popula-
203 tion size is larger than the drift-selection barrier ($1/4s_{LA} = 250$ here; Kimura and Ohta,
204 1971). This prediction would also hold in the presence of migration, as migrants, which
205 are less fit than individuals from the local population, will be more rapidly selected
206 against in large populations, reducing their chance of homogenizing diverging popu-
207 lations. In the allopatric neutral scenario, and in a purely theoretical situation without
208 within-population purifying selection against incompatibilities, population size would
209 not have an effect on the duration of speciation (figure 2B, dashed blue curve). When
210 within-population incompatibilities are selected against, however, this purifying selec-
211 tion is more efficient in larger populations, slowing down both the accumulation of
212 incompatibilities and speciation (figure 2B & E, solid blue curve). After a given size
213 threshold, the relationship between population size and speciation duration is thus posi-
214 tive. The same holds with migration, except that, when populations decrease below
215 a certain size threshold, speciation duration increases as populations become smaller,
216 since the probability of a given migrant genotype fixing and homogenizing populations
217 is higher in smaller populations (figure 2B, solid pink curve). Importantly, these re-
218 sults show that faster speciation in larger populations can occur under both ecological
219 and non-ecological speciation, but faster speciation in smaller populations occurs only
220 under non-ecological speciation (with or without gene flow). We additionally find that
221 speciation duration is primarily influenced by the size of the smallest of the two sepa-
222 rate populations in non-ecological speciation, whereas it is primarily influenced by the
223 size of the largest one in ecological speciation (figure S4). This makes intuitive sense,
224 as mutations causing reproductive incompatibilities accumulate faster in the smallest
225 of the two populations if they are not adaptive, and in the largest population if they are
226 adaptive.

227 The incompatibility threshold K , i.e., the number of genetic differences at which indi-
228 viduals can no longer interbreed, reflects the genomic architecture of speciation, rang-
229 ing from few large-effect genes (or structural genomic changes) for small K to many
230 small-effect genes spread across the genome for large K . In a purely theoretical sce-
231 nario without within-population purifying selection against incompatibilities, the dura-
232 tion of speciation increases with K , as a larger number of fixed differences is required
233 for speciation to occur (figure 2C, dashed curves). It is also largely the case when
234 incompatibilities are selected against, except for very small values of K , where specia-
235 tion duration increases (figure 2C, solid curves). This is due to the extreme purifying
236 selection acting on genomic changes with large effects, which hampers their fixation
237 (see the reduced coefficient of fixation R in figure S1 for small K) and slows down specia-
238 tion, despite the small number of fixed differences required for speciation to occur;
239 this effect is reflected by a reduced polymorphism in the population (figure 2F). Hence,
240 the more polygenic speciation is, the more time it is expected to take, except in cases
241 of very large effect genomic changes that have a hard time fixing.

242 Mutation rates, population sizes, and the genomic architecture of speciation influence
243 not only speciation duration, but also the timing and shape of the grey zone, i.e., when
244 and at which pace the between-population compatibility declines. We fitted a decreasing
245 sigmoid to the between-population compatibility $w_b(t)$ (dashed line in [figure 1](#))
246 and analyzed its timing t_{GZ} (inflection point) and slope r_{GZ} ([Materials and methods](#)).
247 The results mirror the dependencies found for speciation duration, with earlier/faster
248 drops in between-population compatibility corresponding to shorter speciation durations,
249 while later/slower drops correspond to longer speciation durations (see the concordance
250 between [figure 3A](#) and [figure 2A-C](#)). These analyses further show that the
251 only scenario with an early drop in between-population compatibility is the scenario
252 with local adaptation when populations are large ([figure 3A2](#), in green), and that migration
253 enforces a slow compatibility drop ([figure 3A2](#), in red). Hence, we expect –
254 most often – a substantial lag between the time when populations split and the time
255 when compatibility between them drops, and fuzzy species boundaries that extend for
256 a long period of time as soon as there is between-population gene flow.

257 Examining the shape of the grey zone as a function of the net divergence (defined as
258 the difference between divergence and polymorphism, $D_a := D_b - D_w$) instead of time,
259 as is typically done in genomic-based empirical analyses of the grey zone of speciation
260 (Roux et al., 2016; Monnet et al., 2025), reveals a completely different picture
261 ([figure 3B](#)). In this case, the timing and slope of the grey zone are consistent across
262 scenarios. Thus, local adaptation and gene flow have a minor influence on the shape
263 of the genomic grey zone. Population size, mutation rate and genomic architecture all
264 affect the timing of the drop, while its slope is primarily influenced by population size,
265 with a more marginal effect of the other parameters. Earlier drops, reflecting decreases
266 in compatibility at lower levels of divergence, occur under more frequent mutations,
267 larger population sizes, and larger mutational effects. This is explained by the larger
268 polymorphism of the populations at equal level of divergence under these conditions
269 (see [figure S2](#)), which makes it more likely to find pairs of individuals that are not
270 compatible between the two populations. Populations of small size or small mutation
271 rate, on the contrary, maintain a high level of compatibility while accumulating net
272 divergence, until the latter reaches K and compatibility drops sharply. Slower drops,
273 reflecting a fuzzier level of divergence at which speciation occurs, take place in larger
274 populations, and, with a smaller effect, under more frequent mutations and smaller
275 mutational effects. The difference between the genomic and temporal grey zones ([figure](#)
276 [3B](#) versus [figure 3A](#)) show that viewing genomic divergence as a proxy for time
277 since the separation of populations in empirical grey zone curves can lead to biased
278 interpretations. Our theoretical results on the genomic grey zone should help interpreting
279 the shape of the grey zone obtained from empirical data (Roux et al., 2016; Monnet
280 et al., 2025). For example, Monnet et al. (2025) found that reproductive isolation is
281 achieved at a smaller level of divergence in plants compared to animals. They hypothesized
282 that this difference could be explained by reproductive isolation occurring with
283 fewer reproductive barriers in plants. We show here that the number of loci required to
284 achieve reproductive isolation (K) indeed has a large effect on the timing of the grey
285 zone ([figure 3B](#)). Our results suggest that larger population sizes and/or higher mutation
286 rates in plants compared to animals could also explain the pattern. Differences in

287 gene flow or local adaptation, to the contrary, are not expected to impact the shape of
 288 the genomic grey zone.

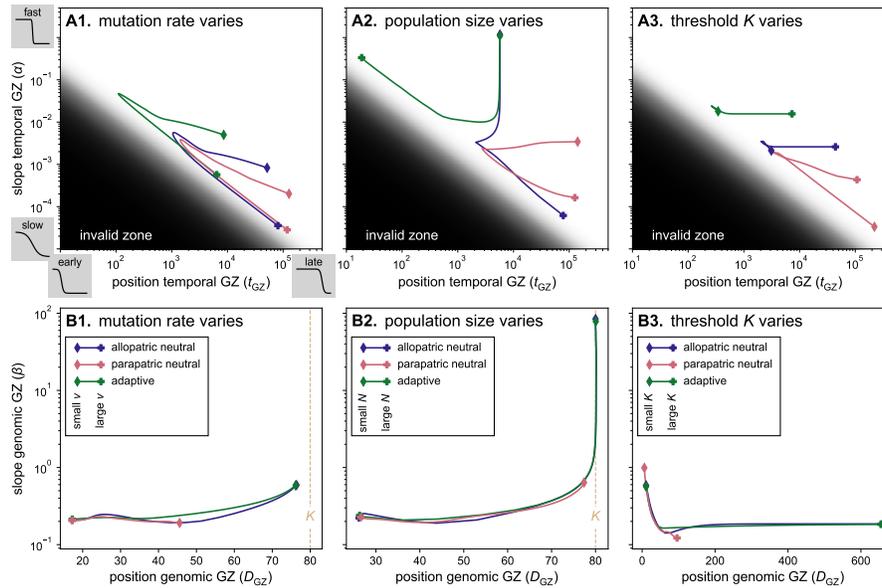


Figure 3. Characteristics of the temporal (A) and the genomic (B) grey zone (GZ) of speciation, with varying mutation rate (1), population size (2) and incompatibility threshold (3). The temporal GZ (resp. genomic GZ) are characterized by a slope α (resp. β) and a position t_{GZ} (resp. D_{GZ}), see [Materials and methods](#). On each line, one parameter changes (increasing from \blacklozenge to \blackplus), each line corresponds to one scenario of speciation. On the upper row, the dark zone of the graphics indicates mathematically unreachable regions; typically early speciation with small slope would require that initial compatibility w_b significantly different from 1 (the middle of the border correspond to $w_b = 0.75$).

289 Our theoretical predictions indicated that a positive correlation between population size
 290 and the duration of speciation would be indicative of non-ecological speciation and
 291 that, in this case, population duration should be primarily influenced by the size of
 292 the smallest population. We evaluated these correlations using genomic-based demo-
 293 graphic inferences for 196 pairs of plant species for which speciation was inferred to
 294 be complete (i.e., no gene flow at present; Monnet et al., 2025). We used posterior
 295 distribution estimates of speciation duration and of the sizes of the ancestral and two
 296 descendant populations, and compared the correlations with a null model to account
 297 for the structure in pairs of the data (see [subsection 4.4](#)). The correlation between speciation
 298 duration and population size was significantly positive ($p < 0.05$) in 98.8% of
 299 the posterior samples when considering the small descendant population (25% in the
 300 null model), 59.5% of the samples when considering the larger descendant population
 301 (22.2% in the null model), and none of the samples when considering the ancestral
 302 population (22% in the null model; [figure 4](#)). These results suggest that non-ecological
 303 speciation dominates in these plant species pairs. The absence of correlation between

304 the ancestral population size and the duration of speciation may seem surprising at
 305 first sight, since population size typically influences the initial polymorphism within-
 306 and between-populations. However, the positive correlation between speciation du-
 307 ration and descendant population sizes suggests that we are in the large-population
 308 regime (figure 2B, right-most part) where polymorphism is expected to be independent
 309 of population size (figure 2E, right-most part). The empirical correlations are there-
 310 fore strikingly consistent with expectations under non-ecological speciation and large
 311 population sizes.

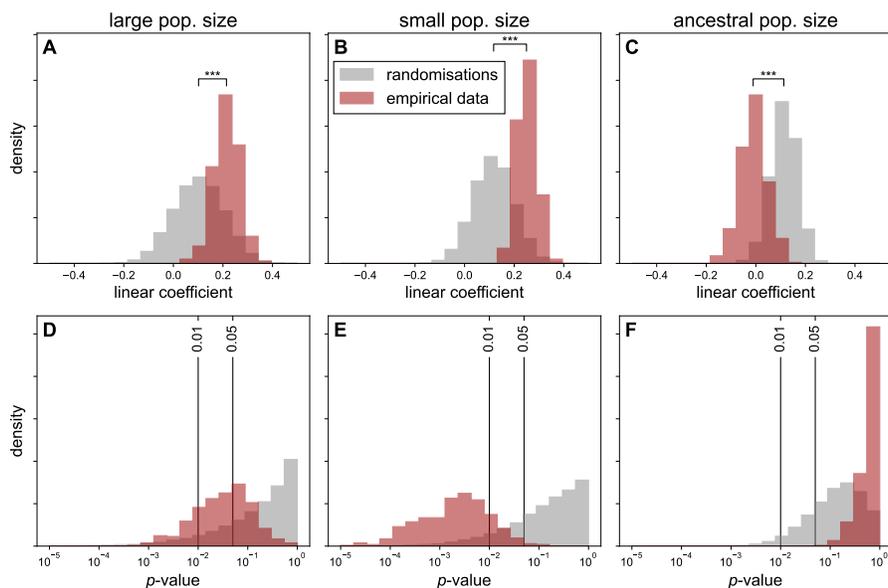


Figure 4. Distribution of linear coefficients (A, B, C) and p -values (D, E, F) for the correlation between speciation duration and population size across posterior samples (red) and randomisations (grey). We fitted the model: $\log(\text{spec. time}) \sim \log N_{\text{large}} + \log N_{\text{small}} + \log N_{\text{ancestral}}$ to 400 posterior samples from demographic parameters inferred on 196 pairs of plants, compared with 100 randomisations between pairs of species for each sample. On the upper row, the difference between the linear coefficient from the randomisations and the empirical data is assessed with a paired t -test (see detailed results in table S1).

312 **3 Conclusion**

313 We investigated how mutation rates, population sizes, the genomic architecture of spe-
 314 ciation, selection, and migration shape the tempo of speciation in a multigenic, corre-
 315 lated adaptive landscape. As expected, higher mutation rates accelerated speciation,
 316 consistent with the evolutionary speed hypothesis, whereas speciation proceeded more
 317 slowly when mutations were neutral than when they were adaptive and in the pres-
 318 ence of gene flow. We also uncovered several unexpected patterns, including a rever-

319 sal of the relationship between mutation rates and speciation duration expected under
320 the evolutionary speed hypothesis at very high mutation rates, and the insensitivity
321 of the genomic grey-zone of speciation to gene flow. Our theoretical and empirical
322 results linking speciation duration to population size revive Mayr's verbal theory of
323 genetic revolution. Specifically, speciation can occur rapidly in small populations in
324 the absence of selection, due to the increased possibility of fixing incompatibilities,
325 and empirical data from 196 plant species pairs are consistent with this non-ecological
326 speciation scenario.

327 However, smaller populations are also more vulnerable to extinction and may col-
328 lapse before speciation is completed. This trade-off underscores the need, in the fu-
329 ture, to jointly consider speciation duration and lineage persistence to understand how
330 microevolutionary processes translate into macroevolutionary speciation rates. The
331 framework developed here provides a foundation for the development of such integra-
332 tive analyses.

333 4 Materials and methods

334 4.1 Predictions under the holey adaptive landscape model

335 Following Gavrillets (1999), we use a system of ordinary differential equations (ODE)
 336 to approximate the mean polymorphism D_w (genetic distance between pairs of indi-
 337 viduals in a population), the mean divergence D_b (genetic distance between pairs of
 338 individuals from two separate populations), and the number of distinct substitutions
 339 (fixed alleles) between the two populations k . Under the assumptions of linkage equi-
 340 librium and rare alleles, the within- and between populations mean compatibilities are
 341 given by (Gavrillets, 1999, eq. 12 and 15):

$$342 \quad w_w = \frac{\Gamma(K+1, D_w)}{\Gamma(K+1)} \quad \text{and} \quad w_b = \frac{\Gamma(K+1-k, D_b-k)}{\Gamma(K+1-k)}$$

343 with $\Gamma(\cdot, \cdot)$ and $\Gamma(\cdot)$ the upper incomplete gamma function and complete gamma func-
 344 tion, respectively. The effect of purifying selection against incompatibilities is captured
 345 by a selection coefficient that depends on the level of within-population polymorphism
 346 (Gavrillets, 1999, eq. 11):

$$347 \quad s = \frac{e^{-D_w} D_w^K}{\Gamma(K+1, D_w)}.$$

348 In the absence of local adaptation, the dynamics follow the system of equations:

$$349 \quad \begin{cases} \frac{dD_w}{dt} = -sD_w + 2\nu - \frac{D_w}{N} + 2m_e(D_b - D_w), \\ \frac{dD_b}{dt} = -s(D_b - k) + 2\nu + 2m_e(D_w - D_b), \\ \frac{dk}{dt} = 2\nu R 2^{-2Nm_e} - 2km_e R \left(\frac{e}{2}\right)^{2Nm_e}. \end{cases} \quad (2)$$

350 where N is the size of each population, ν is the number of mutations per individual
 351 per generation, denoted thereafter as mutation rate, $m_e := m \times w_b/w_w$ is the effective
 352 migration rate, m the migration rate (probability for an individual to have migrated
 353 from the other population at the previous generation), s is the selection coefficient
 354 defined above, and R a fixation coefficient capturing the relative speed of allele fixation
 355 compared to a hypothetical case without incompatibility. The expression of R is given
 356 by (Gavrillets, 1999, eq. 13b):

$$357 \quad R = \frac{2e^{-S} \sqrt{S}}{\sqrt{\pi} \operatorname{erf}(\sqrt{S})}, \quad (3)$$

358 with erf the error function and $S := Ns/2$.

359 In the particular case without migration ($m = m_e = 0$), and under the assumption that
 360 D_w reached an equilibrium (so R is constant), we can obtain an analytical solution for
 361 the duration of speciation by solving the equation for k , which reduces to $\frac{dk}{dt} = 2\nu R$.
 362 When the ancestral population splits, individuals are distributed randomly in the two

363 populations, and it is unlikely that an allele is absent in one population and fixed in
364 the other. Therefore, if $t = 0$ represents the splitting time, the initial number of distinct
365 substitutions is $k(0) = 0$, and thus $k(t) = 2\nu Rt$. Considering that speciation is achieved
366 when $k(t) = K$, the duration of speciation is given by $\tau = \frac{K}{2\nu R}$. If the two populations
367 have different sizes N_1 and N_2 , and therefore different coefficients of fixation R_1 and
368 R_2 , the equation becomes [equation 1](#).

369 To analyse the dynamics in the neutral scenarios, we numerically solved the system
370 given in [equation 2](#). In the presence of migration, speciation is not systematic. In
371 this case, predicting whether speciation will occur or not based on the ODEs can be
372 challenging and time consuming. We designed an alternative, fast approach, based on
373 the bifurcation diagram for the equilibriums of this system. The detailed method is
374 provided in [supplementary text S1](#).

375 To analyse the dynamics in the allopatric scenario with local adaptation, we used the
376 system of ODEs:

$$377 \begin{cases} \frac{dD_w}{dt} = -sD_w + 2\nu - \frac{D_w}{N} \\ \frac{dD_b}{dt} = 2\nu R_{LA}, \\ \frac{dk}{dt} = 2\nu R_{LA}. \end{cases} \quad (4)$$

378 with R_{LA} the coefficient of fixation in the case of local adaptation, given by (Gavrilets,
379 [1999](#), eq 19):

$$380 R_{LA} = \frac{4e^{-S(1-\alpha)^2} \sqrt{S}}{\sqrt{\pi} [\operatorname{erf}(\sqrt{S}(1+\alpha)) + \operatorname{erf}(\sqrt{S}(1-\alpha))]}$$

381 with $\alpha = s_{LA}/s$ denoting the importance of selection due to adaptation compared to
382 the selection due to incompatibility. The equation for D_w is an approximation that
383 ignores the effect of local adaptation (the equation is the same as in the neutral sce-
384 nario with $m_e = 0$) and is justified by the fact that increases in polymorphism due to
385 positive selection when alleles are in low frequency are counter-balanced by decreases
386 when they are in high frequency. The equation for D_b is an asymptotic approximation
387 (Gavrilets, [1999](#), eq. 13a); this equation shows that R_{LA} captures the acceleration of
388 between-population genetic divergence induced by the divergent selection in the two
389 populations.

390 We used Python 3.13 and the module `scipy` (Virtanen et al., [2020](#)) to numerically
391 solve the ODEs. We provide a tool to resolve the equations directly in command line,
392 available on github.com/pierre-veron/HoleyAdaptSpeciation. In order to speed up con-
393 vergence, we used some numerical approximations for some of the calculations, such
394 as the fixation coefficient. Our implementation provides the option to use either the
395 exact form or the approximation.

396 4.2 Simulations

397 The predictions above rely on several simplifications: rare allele assumption, linkage
398 equilibrium, approximation of stochastic processes by deterministic equations, and ad-
399 ditional approximations in the scenario with local adaptation. To verify the validity of
400 these simplifications, we compared the output of the equations with stochastic simula-
401 tions (supplementary text S2). We provide a tool to run the simulations in command
402 line, available on github.com/pierre-veron/HoleyAdaptSpeciation. Without migration,
403 we find that the equations predict well the dynamics observed under the simulations, in
404 particular when we added recombination to the simulations to approximate the linkage
405 equilibrium assumption (figures S7 to S10, configurations A2 and B2). In the neutral
406 allopatric scenario, even simulations without recombination matched the predictions
407 well (condition A1). In the presence of migration to the contrary (conditions C to H),
408 outcomes are more stochastic, and some simulations lead to speciation while it is not
409 expected based on our equations, especially when the migration rate is high and popu-
410 lations are small ($N = 100$ for each population) (conditions G and H). This is consistent
411 with the deterministic approximation, which is not valid when populations are small.

412 4.3 Analyses of the grey zone of speciation

413 The grey zone (GZ) of speciation is the period before speciation during which the two
414 populations are not fully compatible nor fully incompatible. We characterized the GZ
415 by analyzing the between-population compatibility w_b as a function of either (i) time
416 (the temporal GZ) or (ii) net divergence $D_a := D_b - D_w$ (the genomic GZ). The latter
417 corresponds to the GZ obtained from empirical genomic data (e.g. Roux et al., 2016;
418 Monnet et al., 2025). We fitted a decreasing sigmoid to these curves. The temporal
419 GZ is thus characterized by $w_b(t) = 1/(1 + e^{\alpha(t-t_{GZ})})$, with α the slope, and t_{GZ} the
420 timing, or position, of the temporal GZ. The genomic GZ is similarly characterized by
421 $w_b(D_a) = 1/(1 + e^{\beta(D_a - D_{GZ})})$ with β the slope, and D_{GZ} the position, of the genomic
422 GZ. We performed the fit using the function `curve_fit` from the package `scipy` and
423 we checked that the sigmoid was a good approximation (coefficient of determination
424 $R^2 > 0.971$).

425 4.4 Empirical link between population size and the duration of 426 speciation

427 We used genomic estimates of various population genetic parameters from Monnet et
428 al. (2025). These were obtained by applying the Demographic Inference with Linked
429 Selection (DILS) model (Fraïsse et al., 2021) to 280 pairs of plant species (118 species
430 or populations from 25 genera). DILS outputs a best supported model among ancient
431 migration (AM), secondary contact (SC) and isolation with migration (IM); given our
432 goal to retrieve estimates of speciation duration, we selected the 196 species pairs for
433 which speciation was inferred to be complete, i.e., the AM model (see figure S3). We
434 retrieved posterior distributions of the time of the split (T_{split}) and the time of cessation
435 of gene flow (T_{AM}), and used their difference as our estimate of speciation duration
436 (T_{spec}). We also retrieved the posterior distributions of the sizes of the ancestral (N_a)

437 and two descendant populations (N_{large} for the largest, and N_{small} for the smallest). We
438 tested the relationship between the duration of speciation and the population sizes by
439 fitting an ordinary least square (OLS) regression: $\log T_{\text{spec}} \sim \log N_{\text{large}} + \log N_{\text{small}} +$
440 $\log N_a$ and repeated this over 400 samples.

441 The species pairs we used are possible pairs within each of 25 plant genera (Monnet
442 et al., 2025). Individual species therefore appear in multiple pairs, generating a com-
443 plex pattern of non-independence in the data, structured per genus. To account for this
444 non-independence, we repeated the analysis after randomly permuting species identi-
445 ties within each genus. A potential correlation obtained for these permuted data cannot
446 be explained by a true correlation with population sizes, but is rather due to repeated
447 patterns in the data structure (see figure S3). Comparisons between empirical correla-
448 tions and null model correlations (i.e., those obtained after random permutations) allow
449 for the detection of “true correlations” not due to the way the data are organised. For
450 each empirical sampled, we fitted an OLS on 100 randomised data and compared the
451 obtained linear coefficients using a paired t -test (see table S1).

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456 [math.cnrs.fr/](https://impt.math.cnrs.fr/)).

457 **6 Data availability**

458 All scripts and data used in this analysis are available on Zenodo: [https://doi.org/10.](https://doi.org/10.5281/zenodo.18683265)
459 [5281/zenodo.18683265](https://doi.org/10.5281/zenodo.18683265). A GitHub repository containing the functions to run HAL
460 predictions and simulations in command line is also available: [https://github.com/](https://github.com/pierre-veron/HoleyAdaptSpeciation/)
461 [pierre-veron/HoleyAdaptSpeciation/](https://github.com/pierre-veron/HoleyAdaptSpeciation/).

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598 **Supplementary materials**

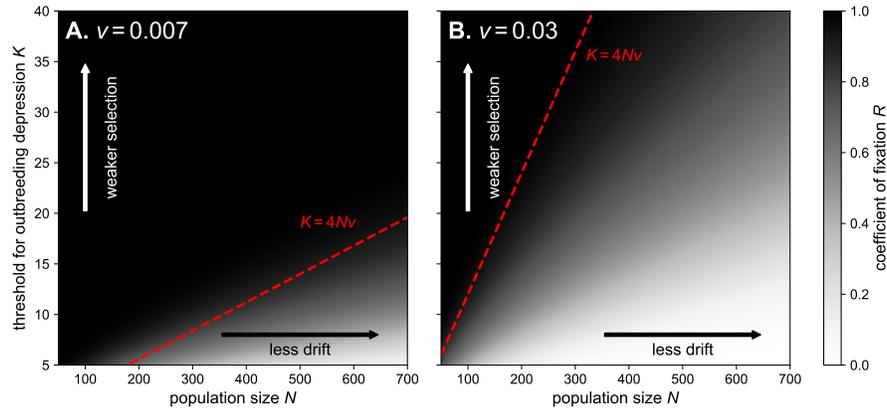
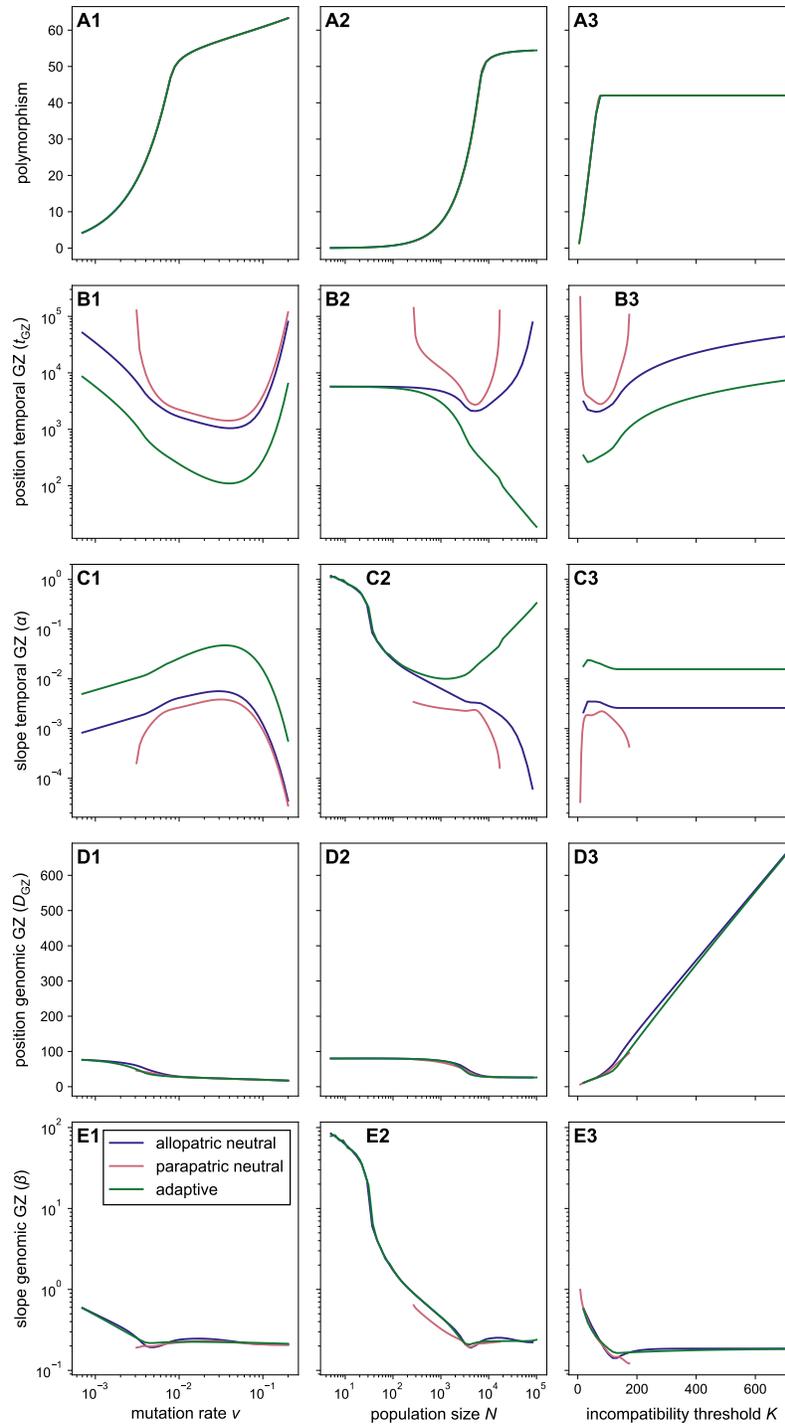


Figure S1. Values of the coefficient of fixation R (representing the relative speed of fixation of mutations compared to a case without purifying selection against incompatibilities, see [equation 3](#)) as a function of the population size N and the incompatibility threshold K , for two values of mutation rate v (**A** and **B**). The dotted red line shows the rule of thumb proposed by Gavrilets (1999) for the limiting conditions where purifying selection against incompatibilities can be neglected ($R \approx 1$).



599

Figure S2. Details of the characteristics of the temporal and genomic grey zones (GZ) of speciation as a function of the mutation rate (1), the population size (2) and the incompatibility threshold (3). **A** shows the mean polymorphism, **B** and **C** show the position and slope of the temporal GZ and **D** and **E** show the position and the slope of the genomic GZ. Each line corresponds to a scenario of speciation.

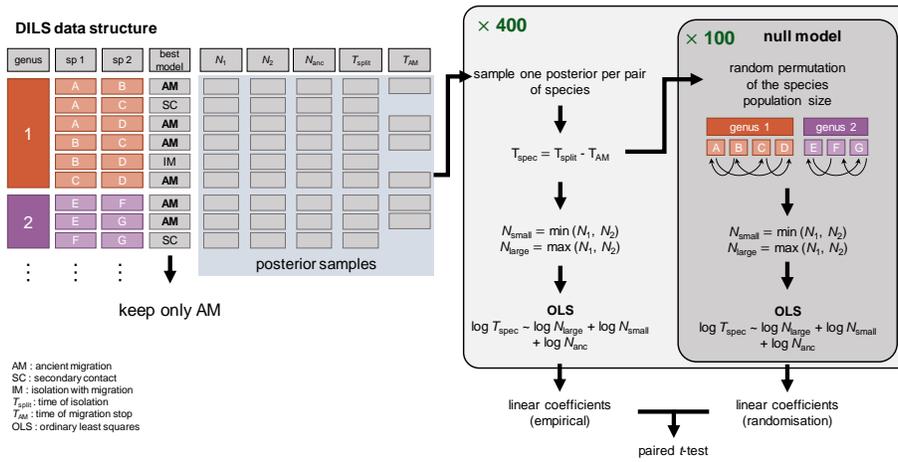


Figure S3. Illustration of the paired structure of the data obtained from DILS, and of the statistical approach used to test the significance of the regression between speciation duration and population sizes . We permuted species within each plant genus to construct a null model.

Table S1. Details of the paired *t*-test between the linear coefficient of the OLS regression on the empirical versus randomized data ($n = 400$ samples).

variable	linear coefficient		<i>t</i>	$\log_{10} p$ -value
	empirical	randomised		
large pop. size	0.214 ± 0.053	0.102 ± 0.102	46.2	-161
small pop. size	0.250 ± 0.038	0.117 ± 0.083	78.6	-244
ancestral pop. size	-0.011 ± 0.054	0.113 ± 0.053	-85.1	-257

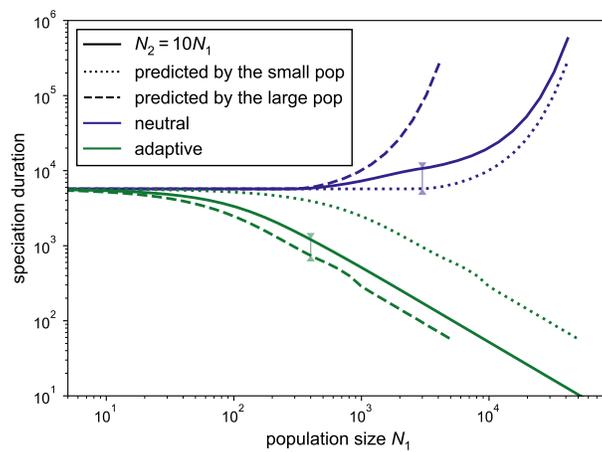


Figure S4. Duration of speciation under the holey adaptive landscape (HAL) allopatric models, in the case of asymmetric populations (solid lines), as a function of N_1 with $N_2 = 10N_1$, compared to the prediction with two populations with size N_1 (dotted line) and two populations with size N_2 (dashed line). In the absence of local adaptation, speciation duration is primarily driven and best predicted by the size of the smallest population (see the small difference between the solid and dotted blue lines, blue arrow); in the presence of local adaptation, it is instead best predicted by the size of the largest population (green arrow).

Supplementary text S1. Determination of the cases without speciation under the parapatric model

In the presence of migration, there are combinations of parameters for which speciation does not occur. Determining these cases by solving the ODEs in [equation 2](#) is not straightforward and time consuming (in particular, speciation not occurring within the set time frame does not imply it will never occur). We therefore designed an alternative approach based on the analyses of the equilibria of the differential equations. For a given combination of parameters, we computed the equilibria by numerically finding the values D_b , D_w and k such that the time derivatives provided in [equation 2](#) are zero. Depending on the parameter values, there are either two equilibrium conditions (one with large D_b , and one with small D_b) or one equilibrium (the two values coincide, see the bifurcation diagrams on [figure S5](#)). These equilibrium conditions are not necessarily met, as they require a specific combination of polymorphism, divergence and number of substitutions. The dynamics of $D_b(t)$ around these equilibria determine whether speciation occurs or not, with $D_b(t)$ diverging to ∞ when speciation occurs. We tested numerically the stability of the equilibrium conditions by resolving $D_b(t)$ with an initial condition close to the equilibrium value and D_w , k taken equal to the equilibrium values. In the cases with one equilibrium, we found that this equilibrium is unstable ($D_b \rightarrow \infty$ as $t \rightarrow \infty$; [figure S6](#) cases 4 and 5); in the cases with two equilibria ([figure S6](#) cases 1, 2 and 3), the lower equilibrium is stable (D_b converges to this equilibrium) and the higher equilibrium is unstable ($D_b \rightarrow \infty$). In practice, the initial divergence $D_b(0)$ is small compared to the equilibrium values, such that condition 3 is not met. Hence, in the parameter space with two equilibria, D_b converges to the lower, stable equilibrium and speciation does not occur. In the parameter space with one equilibrium, to the contrary, speciation occurs and D_b increases without limit. The bifurcation points shown on [figure S5](#) (dotted lines) can thus be used to determine the parameter values at which there is a transition between scenarios with or without speciation (as in [figure 2](#)).

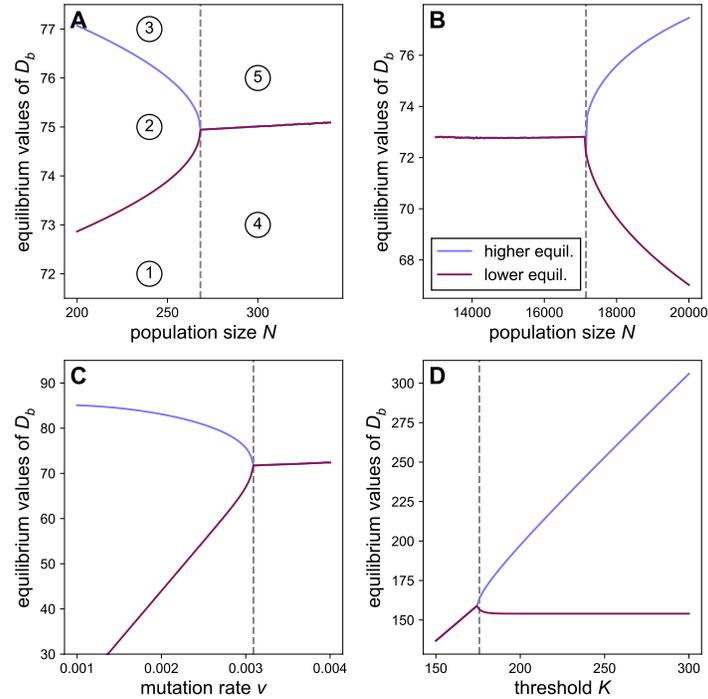


Figure S5. Bifurcation diagrams showing the equilibrium value(s) for D_b as a function of the population size (**A**: low values, **B**: high values), the mutation rate ν (**C**) and the threshold of incompatibility K (**D**). When one parameter varies, the others are kept constant: $\nu = 0.0007$, $N = 6000$, $m = 10^{-4}$ and $K = 80$. The numbers on panel **A** indicate initial conditions for which the dynamics of D_b are represented on [figure S6](#).

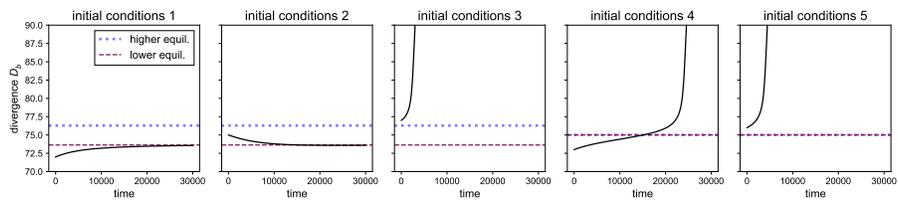


Figure S6. Dynamics of D_b when its initial value is close to the equilibrium and smaller than (1, 4), higher than (3, 5), or between (2) the equilibrium values, corresponding to 5 different regions of the bifurcation diagrams, indicated on [figure S5A](#). The dotted lines indicate the equilibrium value(s). In each case the initial values of D_b and k are the equilibrium values.

Supplementary text S2. Comparison of the HAL deterministic model with simulations

The equations we used to investigate speciation under the HAL model rely on several simplifications, in particular a deterministic approximation (stochastic processes are modelled by ordinary differential equations), an assumption of linkage equilibrium (the allele frequency at a locus is assumed to be independent of the allele frequency at the other loci), a rare allele approximation (the frequency of alleles at polymorphic sites is assumed to be close to 0 or 1 most of the time), and additional approximations detailed in [Materials and methods](#) in the case with local adaptation. In reality, the stochasticity around the equilibrium can play a role in the duration of speciation, especially in small populations where the random reproduction of a small number of individuals can have a strong effect on the dynamics of the whole population. This is exacerbated in the case with migration, as the few individuals who migrate may or may not be compatible with individuals from the receiving population, with important consequences for the speciation process. Furthermore, the assumption of linkage equilibrium is not necessarily met, for example if the recombination rate is low. Gavrillets (1999) performed simulations without recombination and interpreted observed discrepancies between predictions and simulations as the effect of linked loci. We implemented here stochastic individual-based simulations that can account for recombination. To assess the conditions under which the equations are valid and the effect of the linkage equilibrium assumption, we compared the outcomes of the equations with simulations including (or not) recombination. Our simulation process is detailed in [figure S7](#). [figure S8](#) and [figure S9](#) report the results of the comparison with “intermediate” ($N = 600$) and small ($N = 200$) population sizes, which are then summarized in [figure S10](#). With larger population sizes, we expect predictions to be closer to simulations as the deterministic approximation is more valid.

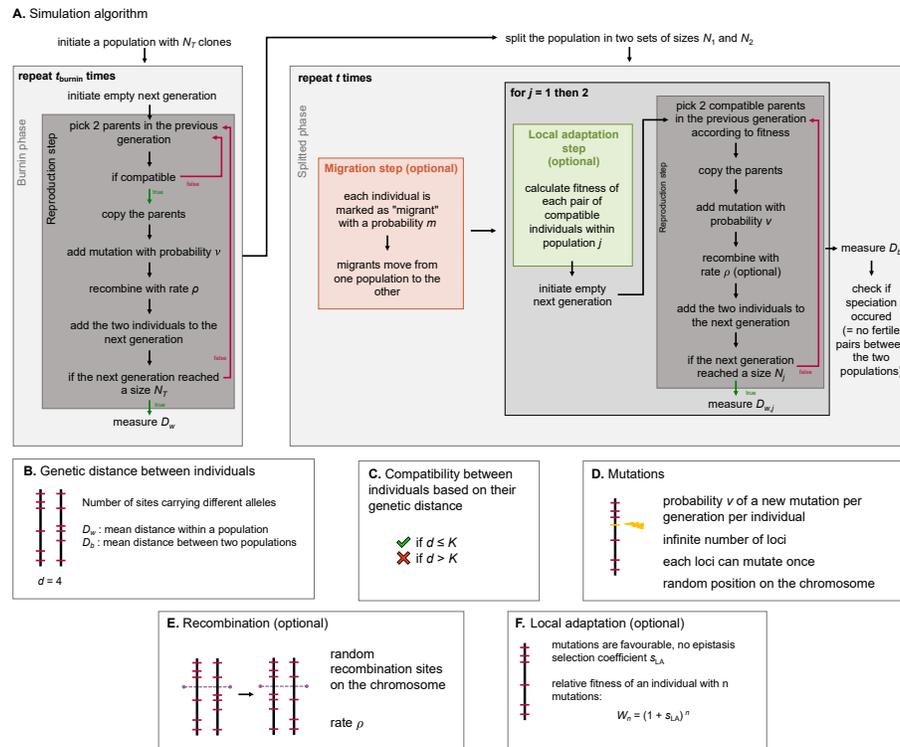
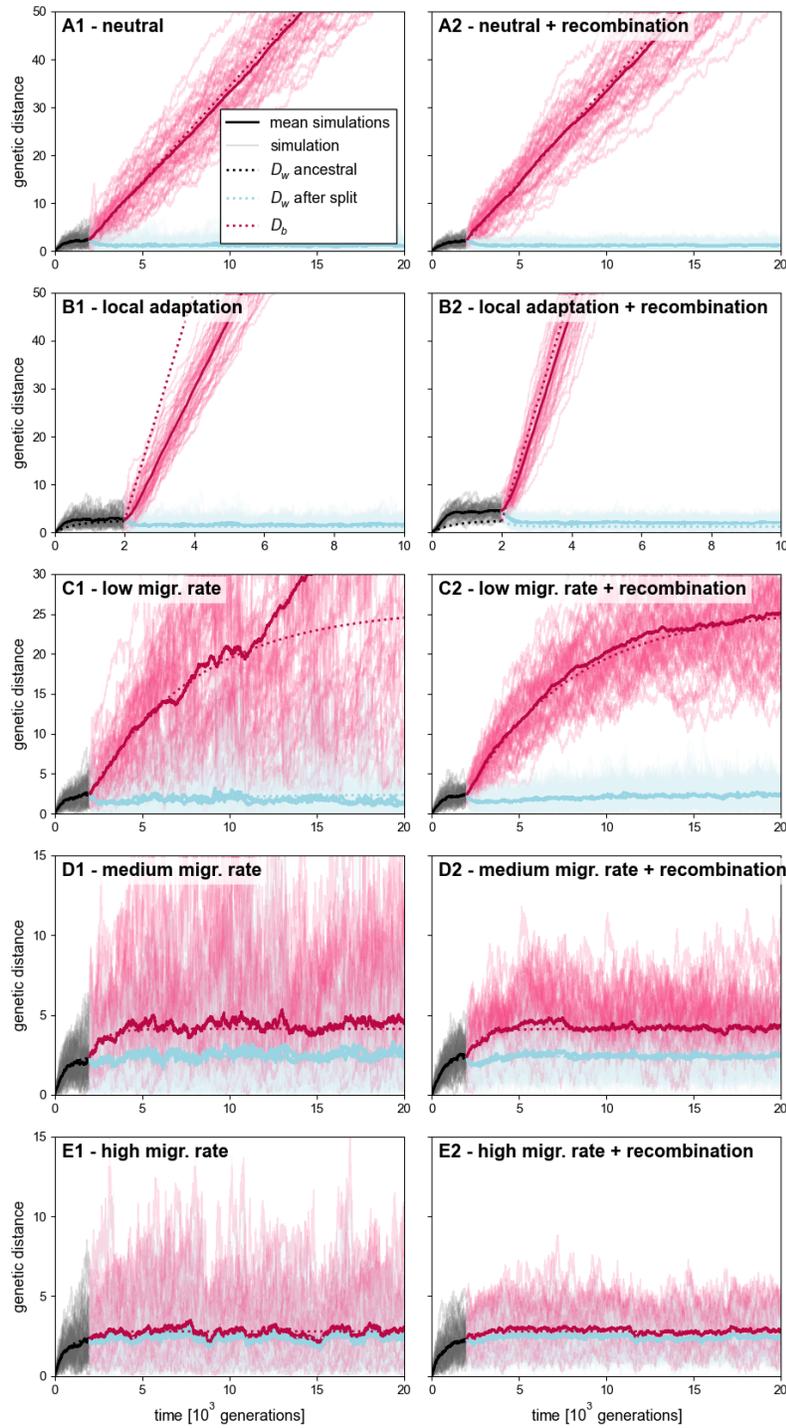


Figure S7. Illustration of the simulation algorithm (**A**) and the several components of the process (**B** to **F**). **A:** A burnin phase is used to allow the ancestral population to reach an equilibrium polymorphism. After this phase, the populations are splitted randomly. **B:** The distance between individuals is calculated as the number of sites carrying different alleles. **C:** The HAL model assumes that individuals can reproduce if their genetic distance is smaller than the incompatibility threshold. **D:** At each generation, each individual can mutate with probability μ . In this case, a site that never carried a mutation is altered. A gene cannot undergo the same mutation twice independently. **E:** For simulations with recombination, a part of the chromosome is randomly exchanged between the two offsprings of a pair of parents. **F:** For simulations with local adaptation, the fitness of the each compatible pair of individuals within each populations is calculated and is used to sample the parents during the reproduction step.



600

Figure S8. Dynamics of polymorphism and divergence under the holey adaptive landscape model obtained with the stochastic simulations (50 simulations for each configuration, solid thin lines ; the solid thick lines show the mean of the simulations) and with the deterministic approximations (dotted lines). The first column shows simulations without recombination and the second column shows simulations with recombination (recombination rate $\rho = 2$). The first row (**A**) corresponds to the allopatric neutral model, the second row (**B**) corresponds to the allopatric model with local adaptation ($s_{LA} = 0.005$) and the last third rows (**C**, **D**, **E**) correspond to the model with low, medium and high migration (migration rate $m = 8.5 \times 10^{-5}$, 0.001155 and 0.00498 respectively). Parameters of the model: $\nu = 0.002$, $N_1 = N_2 = 300$, $K = 40$, burnin time 2000 generations.

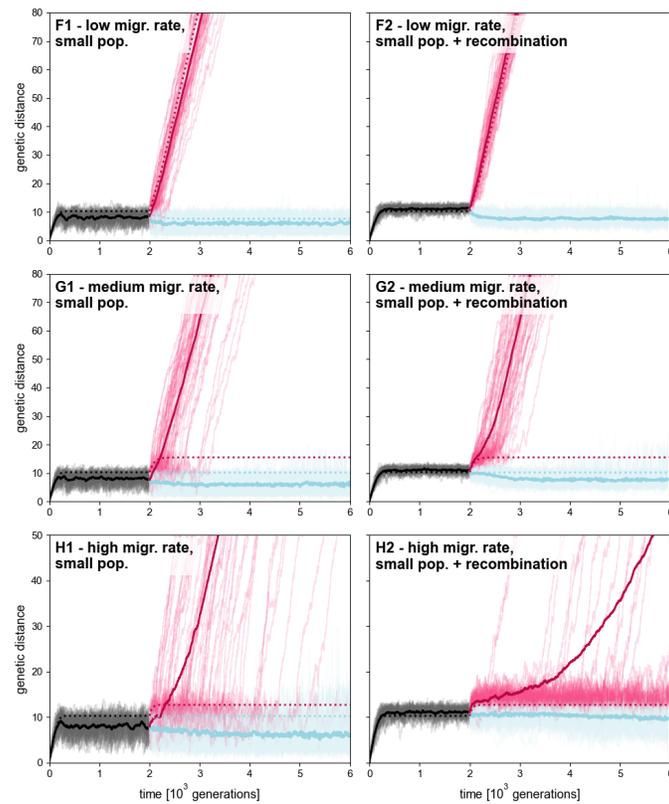


Figure S9. Same as [figure S8](#) but with smaller populations : $\nu = 0.0384$, $N_1 = N_2 = 100$, $K = 20$ and migrations rates 0.001, 0.005 and 0.01 for rows **F**, **G** and **H** respectively.

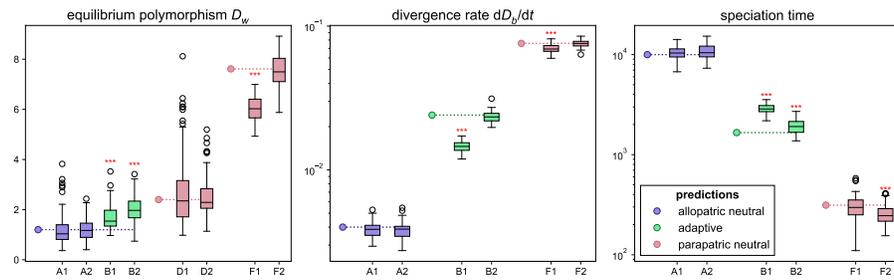


Figure S10. Comparison of the dynamics of the simulations and the predictions, based on (left) the equilibrium polymorphism D_w , (middle) the slope of $D_b(t)$ and (right) the duration of speciation. The boxplots indicate the distribution of the estimates from 50 simulations for each configuration, and the dot the expected value based on the deterministic prediction. The configurations (A1, A2, B1, B2, D1, D2, F1 and F2) are those of figures S8 and S9. The red asterisks indicate significant differences, assessed with t -test and corrected for multiple testing.